

Minimal Clinically Important Difference for PROMIS Physical Function and Pain Interference in Patients Following Surgical Treatment of Distal Radius Fracture

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Purpose We estimated the minimal clinically important difference (MCID) for the Patient-Reported Outcomes Measurement Information System (PROMIS) Physical Function (PF) and Pain Interference (PI) computer adaptive tests (CATs) following surgical treatment of distal radius fracture (DRF).

Methods Adult patients surgically treated between November 2017 and November 2020 for isolated DRF were identified. Demographic and patient-reported outcome data were extracted from the electronic health record. Outcomes of interest were the PROMIS PF and PI CATs. Inclusion criteria were met if: (1) PROMIS PF and PI scores were available at preoperative and postoperative visits; and (2) a postoperative clinical anchor question asking about overall response to treatment was answered. An anchor-based MCID estimate was determined by calculating the average absolute score change in PROMIS PF and PI for patients who indicated a mild change to the anchor question. A distribution-based MCID estimate was also calculated using the standard error of measurement and effect sizes of change.

Results The changes in PROMIS PF and PI scores were significantly different between patients who gave responses of much change ($n = 73$), mild change ($n = 51$), and no change ($n = 19$) to the clinical anchor question. The average score changes in the mild change group for PROMIS PF and PI were 5.2 (SD, 3.7) and 6.8 (SD, 4.3) points, respectively, representing the anchor-based MCID estimates. The PROMIS PI anchor-based estimate was moderately correlated with the preoperative score ($r = -0.41$), time between visits ($r = -0.39$), and age ($r = 0.30$). The distribution-based MCID estimates were 3.8 (SD, 1.3) and 3.7 (SD, 1.3) points for the PROMIS PF and PI, respectively.

Conclusions The MCIDs were estimated as 5.2 and 6.8 for the PROMIS PF and PI CATs, respectively, following surgery for DRF.

Clinical relevance As reports continue to publish a consistent range of MCID values, researchers can be confident in these values and begin using them across a broader spectrum of conditions treated by hand surgeons. (*J Hand Surg Am.* 2022;47(2):137–144. Copyright © 2022 by the American Society for Surgery of the Hand. All rights reserved.)

Key words Distal radius fracture, MCID, minimal clinically important difference, patient-reported outcomes, PROMIS.

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Received for publication February 27, 2021; accepted in revised form August 19, 2021.

No benefits in any form have been received or will be received related directly or indirectly to the subject of this article.

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0363-5023/22/4702-0004\$36.00/0
<https://doi.org/10.1016/j.jhsa.2021.08.025>

DISTAL RADIUS FRACTURE (DRF) is a common injury in the United States,^{1–4} accounting for 15% to 20% of all fractures treated in the emergency setting.⁵ The overall incidence of DRF among US adults has reportedly increased over the past 2 decades,^{2,6} and consequently represents a growing economic burden.^{7,8} Patient-reported outcome (PRO) measures are being adopted by practitioners to capture the impact of disease on patient health status and determine the effectiveness of treatment. By providing a patient-centered view of current health status, PRO measures can promote shared clinical decision-making and guide expectations following treatment.

Commonly used region- and condition-specific PRO measures in hand surgery include the Disabilities of the Arm, Shoulder, and Hand questionnaire (DASH),⁹ *QuickDASH*,¹⁰ Michigan Hand Outcomes Questionnaire (MHQ),¹¹ Patient-Rated Wrist Evaluation,¹² and Boston Carpal Tunnel Questionnaire (BCTQ).¹³ These questionnaires are validated and reliable measures of hand function. However, barriers exist preventing their full adoption and standardization across hand surgery. Several authors have reported concern regarding their narrow scopes, administrative burdens, and notable ceiling and floor effects.¹⁴ In 2004, the PROMIS was developed by the National Institutes of Health to improve the overall reporting of physical, mental, and social health in a maximally precise and efficient manner.¹⁵ The PROMIS is a general PRO measure that uses item response theory (IRT) and computerized adaptive testing (CAT), which have the benefits of decreased respondent and administrative burdens while maintaining accuracy and reproducibility.^{16,17} When compared to legacy instruments, the PROMIS has performed favorably by demonstrating excellent reliability (ie, reproducibility), responsiveness (ie, ability to detect score changes), one-dimensionality (ie, low unexplained variance), and coverage (ie, minimal ceiling and floor effects).^{18–20}

The PROMIS is increasingly being used to evaluate outcomes following hand/wrist procedures. To understand the clinical implications of postoperative score changes over time, it is important to know the minimal clinically important difference (MCID), which represents the smallest change in an outcome score that a patient may perceive as clinically meaningful (either beneficial or harmful).²¹ The MCID is an intrinsic property of a PRO measure, and variation in estimates of the MCID reflects the responsiveness of the PRO measure to the different conditions studied. MCID calculations are most often

performed with anchor- or distribution-based methods. Anchor-based methods compare changes in PRO scores to an external, independent assessment of clinical change,^{22,23} whereas distribution-based methods use statistical calculations based on the variability in scores from the sampled population to determine the level of change that is beyond random occurrence.²⁴ The anchor-based method is generally regarded as the preferred approach because it is patient-centered; however, both methods have their limitations, and it has therefore been recommended that MCID estimates be based on multiple approaches, ideally resulting in values that converge on a single value or narrow range of values.²⁵

Several studies within the hand literature have reported varying estimates of the true MCID for the PROMIS PF and PI CATs.^{26–30} While there is no consensus regarding the amount of data necessary prior to establishing the true MCID for a PRO measure, there is no doubt that additional data extend confidence in the previously reported values. Therefore, the objective of this study was to estimate the MCID for PROMIS PF and PI CATs in patients surgically treated for DRF using an anchor- and distribution-based approach.

MATERIALS AND METHODS

Institutional review board approval was obtained from the University of Rochester Medical Center to review our longitudinally maintained PROMIS database. Patients who presented to our hand clinic between November 2017 and November 2020 for elective outpatient surgical treatment of isolated DRF were identified using Current Procedural Terminology (CPT) billing codes 25607, 25608, and 25609. Demographic and PRO data were automatically extracted from the electronic health record. Inclusion criteria were patients 18 years of age and older with DRF who were treated with surgery using open reduction and internal fixation. In addition, PROMIS data had to be available at both preoperative (<21 days before surgery) and postoperative (<365 days after surgery) visits. Patients who failed nonoperative management and initially presented beyond 21 days prior to surgery were excluded. The postoperative visit closest to the 6-week follow-up date was chosen for analysis. A clinical anchor question also had to be completed at the same postoperative visit as the PROMIS score selected to be included in the analysis. A flowchart diagram detailing our inclusion and exclusion criteria is illustrated in [Figure 1](#).

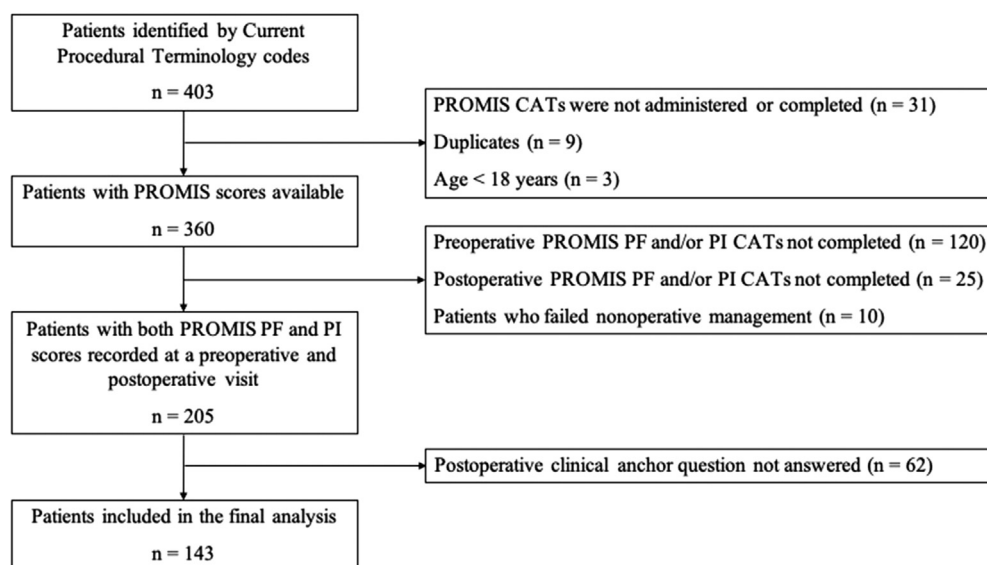


FIGURE 1: Flowchart illustrating inclusion and exclusion criteria.

The PROMIS CATs of interest were PF (versions 1.2 and 2.0) and PI (version 1.1). As part of routine clinical care, these CATs were administered to patients at office visits using an Apple iPad. During the study period, we transitioned from using PROMIS PF version 1.2 to version 2.0; however, scores across these versions can be compared to each other.³⁰ The PROMIS PF CAT measures self-reported physical capability, whereas the PROMIS PI CAT measures the degree to which pain interferes with physical, mental, and social activities.³¹ Higher PROMIS PF scores indicate better physical capability, whereas lower PROMIS PI scores indicate better pain coping. PROMIS scores are reported as a *t* score, which is a standardized score referenced against the general US population.³² PROMIS scores follow a normal distribution, with a mean *t* score of 50 and an SD of 10.^{33,34}

To independently assess patient-perceived clinical changes, the following anchor question was asked to each patient at their postoperative visits: “since your first visit to the provider you are seeing today, has your condition treated by this provider changed?” Responses to the anchor question were limited to a scale from 0 to 5, with 0 meaning “I am seeing the physician for a new problem,” 1 indicating the condition was “much better,” 2 that the condition was “mildly better,” 3 that there was “no change,” 4 that the condition was “mildly worse,” and 5 that the condition was “much worse.”

Statistical analysis

Descriptive statistics were calculated. Continuous variables are reported as means \pm SDs and

categorical variables are presented as frequencies and percentages. Statistical significance was set at a *P* value $< .05$.

Analysis of variance with Tukey’s *post hoc* testing was done for significant pairwise comparisons to determine whether the average absolute changes in PROMIS PF and PI *t* scores were significantly different between visits when patients reported much change (eg, “much better” or “much worse”), mild change (eg, “mildly better” or “mildly worse”), and “no change.” This allowed us to determine whether numerical changes in PROMIS scores reflected a true change as perceived by the patient. The average change in score among patients reporting mild change was chosen as the anchor-based MCID estimate for the respective PROMIS domain.

Bivariate statistical testing was performed to identify patient variables affecting the anchor-based MCID estimate. Among patients who reported mild change since their preoperative visit, Spearman’s correlation coefficients were calculated to determine the association between the change in PROMIS PF and PI scores and the preoperative PROMIS score for the respective PROMIS domain, number of days between visits, age, and body mass index (BMI).

A distribution-based estimate was also calculated using methods originally described by Yost et al.^{35,36} These methods rely on the statistical distribution of PRO data, including standard error of measurement (SEM) and effect size measures. The SEM, which represents the minimal detectable change (MDC) in a scale, reflects the smallest score change likely to be a true change rather than measurement error. In IRT,

each *t* score is associated with a standard error. Therefore, using the previous 3,900 patient encounters at our hand clinic, the MDC was calculated as the average standard error across the PROMIS PF and PI CATs. The MDC values for the PROMIS PF and PI CATs were estimated to be 2.4 and 2.2, respectively. To be clinically relevant, the MCID estimate must exceed the MDC. Next, an effect size filter was applied to the anchor-based MCID *t* scores by taking the change in PROMIS score and dividing by the baseline SD of the sample.^{37,38} Only score differences corresponding to effect sizes between 0.2 and 0.8 were considered in the final distribution-based MCID estimate. Score differences corresponding to effect sizes less than 0.2 (Cohen's small effect size) were deemed unlikely to be clinically important, whereas scores differences corresponding to effect sizes greater than 0.8 (Cohen's large effect size) were deemed unlikely to be minimal.

RESULTS

Baseline patient characteristics are summarized in Table 1. Of the 143 patients that met the inclusion criteria, most were female (*n* = 112; 78.3%), White (*n* = 132; 90.2%), and non-Hispanic (*n* = 119; 83.2%). The average patient age was 56.6 years old (SD, 16.3 years), and most patients were nonsmokers (*n* = 91; 63.6%) and not obese (*n* = 92; 64.3%).

Patients presented to our clinic at a median of 5 days (interquartile range [IQR], 3–6 days) prior to surgery. Preoperative PROMIS PF scores averaged 33.8 (SD, 8.1), with a range of 22.8 to 66.6, and preoperative PROMIS PI scores averaged 65.1 (SD, 7.1), with a range of 38.7 to 80.1.

Patients had a postoperative visit with clinical anchor data recorded at a median of 34 days (IQR, 11–49 days) after surgery. The average postoperative PROMIS PF score was 39.4 (SD, 8.9; range, 23.2–66.0), while the average postoperative PROMIS PI score was 57.5 (SD, 8.1; range, 38.7–77.8).

The changes in PROMIS PF and PI scores between visits were significantly different between patients who reported much change (*n* = 73), mild change (*n* = 51), and no change (*n* = 19) to the clinical anchor question (*P* < .05). Using an anchor-based approach for estimating the MCID, the average changes in PROMIS PF and PI scores for the mild change group were +5.2 (SD, 3.7) and –6.8 (SD, 4.3), respectively (Tables 2 and 3); thus, both MCID anchor-based estimates exceeded the MDC level

TABLE 1. Baseline Patient Characteristics

Characteristics	Values
Total number of patients	143
Age, years (range)	56.6 (18.9–85.9)
Sex, n (%)	
Female	112 (78.3)
Male	31 (21.7)
Race, n (%)	
White	129 (90.2)
Black	6 (4.2)
Other	5 (3.5)
Unknown	3 (2.1)
Ethnicity, n (%)	
Not Hispanic	119 (83.2)
Hispanic	5 (3.5)
Unknown	19 (13.3)
BMI, kg/m ² , n (%)	
<25	55 (38.5)
25–30	37 (25.9)
>30	51 (35.7)
Smoking status, n (%)	
Never	91 (63.6)
Former	33 (23.1)
Current	19 (13.3)
CPT code,* n (%)	
25607	10 (7.0)
25608	51 (35.7)
25609	82 (57.3)

*CPT, Current Procedural Terminology.

TABLE 2. Average Change in PROMIS Physical Function According to the Clinical Anchor Response

Clinical Anchor Response	Δ Physical Function	SD
Much change (<i>n</i> = 73)	+10.2	6.6
Mild change (<i>n</i> = 51)	+5.2*	3.7
No change (<i>n</i> = 19)	1.8	2.3

*MCID estimate.

of instrument error (PROMIS PF MDC, +2.4; PROMIS PI MDC, –2.2).

Among patients who reported a mild change since their preoperative visit, using Spearman's correlation coefficient, the degree of score change for the

TABLE 3. Average Change in PROMIS Pain Interference According to the Clinical Anchor Response

Clinical Anchor Response	Δ Pain Interference	SD
Much change (n = 73)	-12.7	6.8
Mild change (n = 51)	-6.8*	4.3
No change (n = 19)	2.1	3.1

*MCID estimate.

PROMIS PF CAT was not correlated with the preoperative score ($r = -0.19$; $P = .25$), age ($r = -0.23$; $P = .46$), BMI ($r = 0.04$; $P = .79$), nor time between visits ($r = -0.07$; $P = .68$). For the PROMIS PI CAT, no correlation was observed between the degree of score change and BMI ($r = 0.17$; $P = .24$); however, weak to moderate correlations were observed between the degree of score change and the preoperative score ($r = -0.41$; $P = .003$), time between visits ($r = -0.39$; $P = .005$), and age ($r = 0.30$; $P = .04$).

A distribution-based MCID estimate was also calculated for the PROMIS PF and PI CATs. After applying effect size parameters of 0.2 to 0.8 to data from patients who reported a mild change since their preoperative visit, distribution-based MCID values were estimated to be +3.8 (SD, 1.3) and -3.7 (SD, 1.3) for the PROMIS PF and PI, respectively (Tables 4 and 5). These distribution-based MCID estimates also exceed the MDC values of +2.4 for the PROMIS PF and -2.2 for the PROMIS PI.

DISCUSSION

This study estimates, using an anchor- and distribution-based approach, the MCID values for the PROMIS PF and PI CATs in patients surgically treated for DRF. The range of estimated MCID values for the PROMIS PF CAT was +3.8 to +5.2 points, while the range of estimated MCID values for the PROMIS PI CAT was -3.7 to -6.8 points. These estimates were greater than the MDC, and therefore likely reflect a true change rather than measurement error. An important point in looking at MCID and MDC is the direction of change. While the common thought is that the MCID shows improvement, the change can indicate improvement or worsening. Similarly, the direction of the MDC can be up or down and represent an improvement or worsening change (depending on the specific PROMIS domain) that was noticeable. We believe this should be taken into consideration when looking at values, planning

TABLE 4. Average Change in PROMIS Physical Function After Distribution Correction

Clinical Anchor Response	Δ Physical Function	SD
Much change (n = 50)	+13.3	5.7
Mild change (n = 27)	+3.8*	1.3
No change (n = 12)	0.5	0.6

*MCID estimate.

TABLE 5. Average Change in PROMIS Pain Interference After Distribution Correction

Clinical Anchor Response	Δ Pain Interference	SD
Much change (n = 64)	-14.1	5.1
Mild change (n = 23)	-3.7*	1.3
No change (n = 12)	0.4	0.5

*MCID estimate.

prospective studies, and determining power or sample size.

The estimates reported in our study compare favorably to previously reported MCID values for the PROMIS PF and PI CATs in the hand literature. Sandvall et al²⁸ reported a range of MCID values for PROMIS PF between 3.6 and 4.6 points in patients treated nonsurgically for DRF. Similarly, in a cohort of patients treated both surgically and nonsurgically for thumb carpometacarpal arthritis, Lee and Calfee²⁹ reported a range of MCID values for PROMIS PF between 3.5 to 3.9 points. Based on validated region-specific (MHQ) and condition-specific (BCTQ) PRO measures, Bernstein et al²⁷ provided MCID estimates for the PROMIS PF (1.8 to 2.8 points) and PI (-4.1 to -9.7 points) for patients undergoing carpal tunnel release. Alternatively, using a half-SD distribution-based method, Kazmers et al³⁰ calculated MCID values of 4.6 and -3.4 for the PROMIS PF and PI CATs in patients with hand and elbow conditions. Our findings seem comparable, particularly given that we looked at a traumatic injury that likely resulted in an acute decline in function and greater pain interference at the initial evaluation and thus would be expected to show greater improvement following treatment when compared to a condition that the patient had potentially adapted to over time. The fact that our estimates fall within a narrow range of previously reported values strongly suggests that a true

MCID value exists for the PROMIS PF and PI scales. This supports the notion that the MCID is not a fluid value, but rather an intrinsic value of a PRO measure. The extent to which estimates of the MCID vary purely reflects varying degrees of responsiveness of the PRO measure to the different hand conditions investigated.

Prior studies in the literature have suggested that MCID estimates are not influenced by the severity of patients' baseline scores nor the length of time since treatment.^{39–41} However, more recent reports in the orthopaedic literature have suggested that preoperative PROMIS scores can predict postoperative score improvement.^{26,42} Similarly, our study found that higher preoperative PROMIS PI scores were moderately associated with the MCID estimate. These findings reflect the ability of the PROMIS scale to detect greater changes among patients with worse baseline scores. In addition, our MCID estimate for PROMIS PI seemed to be influenced by the length of time between the preoperative and postoperative visits. Consequently, future work may be needed to determine whether our PROMIS PI estimate can be applied to patients outside of the 6-week time frame.

Estimating the MCID for the PROMIS CATs is valuable in understanding the instruments' ability to detect clinically meaningful changes in outcome scores over time. Furthermore, MCID estimates can be used to evaluate treatment effectiveness between 2 cohorts, as well as to sufficiently power a clinical study.⁴³ However, summarized by the concept of responsiveness, MCID estimates for a PRO measure cannot be used blindly across all clinical applications and patient samples. Studies have shown that MCID estimates for PRO measures are dependent on the clinical context in which they were initially calculated.^{44–46} This includes patient characteristics of the sample population (eg, socioeconomic status, patient expectations) and the clinical picture (eg, disease type, intervention). For example, a large estimated MCID value for a certain hand condition may indicate that the PRO measure is not sufficiently sensitive to clinical change for that hand condition; therefore, a different, more responsive measure may be required to enhance the efficiency of a clinical study. Alternatively, in the case of a less responsive PRO measure, a greater number of subjects may be required in order to detect significant differences between groups of patients. Consideration for these factors should be made before applying an MCID estimate in clinical research. Future research can look at the responsiveness of PROMIS for a given condition as the MCID ranges seem to be established for a variety of

traumatic and degenerative conditions in the hand and wrist.

Finally, some authors have proposed that MCID values be used to evaluate clinical changes of individual patients.⁴² However, from our experience as well as that of others, we believe that the best application of MCID values is in the evaluation of PRO scores at the group level.^{17,29} At the individual level, we observed a significant variation in scores among patients reporting their condition as "mildly better" or "mildly worse," and therefore the MCID estimate would likely not be a reliable threshold for individual clinical change following surgical treatment of DRF. For individual patients, following the direction of change seems to be more meaningful.

This study has several limitations. First, our MCID estimates are derived from a patient cohort that presented to a single hand clinic for elective outpatient surgery for DRF. Patients were predominately female, White, and non-Hispanic. In addition, our anchor-based estimate was based on PROMIS data and clinical anchor responses from just 51 patients reporting mild changes. These factors should be considered before applying our MCID estimates to a larger, more diverse patient population. Furthermore, while our MCID estimates fall within the range of previously reported estimates for the PROMIS PF and PI CATs, they may not correlate to all patients with operatively treated DRF, particularly those with open injuries or other associated injuries in the same extremity. Second, anchor-based MCID estimates are prone to recall bias because patients are asked to compare their current, posttreatment health status to their pretreatment state. We strived to use postoperative data from the 6-week follow-up visit; however, for patients without these data available, data from earlier or later postoperative time points were chosen. This was to select a relatively uniform time point when most patients are still in treatment but have recovered enough to detect change. Third, our clinical anchor question asked about overall health status when compared to the pretreatment period. Because our anchor question did not explicitly identify the target PRO measure (ie, PF or PI), it is possible that responses to our anchor did not truly reflect the clinical outcome that the PRO instrument intended to measure. However, we used an anchor question that has been previously reported in the hand literature.^{28,29,47} Lastly, while we determined the MCID for the PROMIS PF and PI CATs, the PROMIS Upper Extremity (UE) CAT is also relevant to hand surgery. At the time the patients in our study were initially treated at our hand clinic, the PROMIS

UE CAT was not routinely collected. As hand surgeons work toward a consensus regarding a primary PRO measure to use, future work may need to estimate the MCID as well as responsiveness to the PROMIS UE CAT in patients surgically treated for DRF.

In conclusion, this study adds to the growing hand literature defining an approximate 3- to 6-point score change on the PROMIS PF and PI CATs as the MCID. As studies continue to report a consistent range of MCID values for patients seeking hand care, researchers can gain confidence in these estimates and begin using them across a broader spectrum of conditions treated by hand surgeons.

REFERENCES

- Karl JW, Olson PR, Rosenwasser MP. The epidemiology of upper extremity fractures in the United States, 2009. *J Orthop Trauma*. 2015;29(8):e242–e244.
- Nellans KW, Kowalski E, Chung KC. The epidemiology of distal radius fractures. *Hand Clin*. 2012;28(2):113–125.
- Baron JA, Karagas M, Barrett J, et al. Basic epidemiology of fractures of the upper and lower limb among Americans over 65 years of age. *Epidemiology*. 1996;7(6):612–618.
- Court-Brown CM, Caesar B. Epidemiology of adult fractures: a review. *Injury*. 2006;37(8):691–697.
- Bengné U, Johnell O. Increasing incidence of forearm fractures. A comparison of epidemiologic patterns 25 years apart. *Acta Orthop Scand*. 1985;56(2):158–160.
- de Putter CE, van Beeck EF, Looman CW, Toet H, Hovius SE, Selles RW. Trends in wrist fractures in children and adolescents, 1997–2009. *J Hand Surg Am*. 2011;36(11):1810–1815.e2.
- Zhong L, Mahmoudi E, Giladi AM, Shauver M, Chung KC, Waljee JF. Utilization of post-acute care following distal radius fracture among Medicare beneficiaries. *J Hand Surg Am*. 2015;40(12):2401–2409.e8.
- Huetteman HE, Zhong L, Chung KC. Cost of surgical treatment for distal radius fractures and the implications of episode-based bundled payments. *J Hand Surg Am*. 2018;43(8):720–730.
- Hudak PL, Amadio PC, Bombardier C. Development of an upper extremity outcome measure: the DASH (disabilities of the arm, shoulder and hand) [corrected]. The Upper Extremity Collaborative Group (UECG). *Am J Ind Med*. 1996;29(6):602–608.
- Gummeson C, Ward MM, Atroshi I. The shortened disabilities of the arm, shoulder and hand questionnaire (*QuickDASH*): validity and reliability based on responses within the full-length DASH. *BMC Musculoskelet Disord*. 2006;7:44.
- Chung KC, Hamill JB, Walters MR, Hayward RA. The Michigan Hand Outcomes Questionnaire (MHQ): assessment of responsiveness to clinical change. *Ann Plast Surg*. 1999;42(6):619–622.
- MacDermid JC, Turgeon T, Richards RS, Beadle M, Roth JH. Patient rating of wrist pain and disability: a reliable and valid measurement tool. *J Orthop Trauma*. 1998;12(8):577–586.
- Leite JC, Jerosch-Herold C, Song F. A systematic review of the psychometric properties of the Boston Carpal Tunnel Questionnaire. *BMC Musculoskelet Disord*. 2006;7:78.
- Hsu JE, Nacke E, Park MJ, Sennett BJ, Huffman GR. The Disabilities of the Arm, Shoulder, and Hand questionnaire in intercollegiate athletes: validity limited by ceiling effect. *J Shoulder Elbow Surg*. 2010;19(3):349–354.
- Brodke DJ, Saltzman CL, Brodke DS. PROMIS for orthopaedic outcomes measurement. *J Am Acad Orthop Surg*. 2016;24(11):744–749.
- Cella D, Yount S, Rothrock N, et al. The Patient-Reported Outcomes Measurement Information System (PROMIS): progress of an NIH Roadmap cooperative group during its first two years. *Med Care*. 2007;45(5):S3–S11.
- Hammert WC, Calfee RP. Understanding PROMIS. *J Hand Surg Am*. 2020;45(7):650–654.
- Döring AC, Nota SP, Hageman MG, Ring DC. Measurement of upper extremity disability using the Patient-Reported Outcomes Measurement Information System. *J Hand Surg Am*. 2014;39(6):1160–1165.
- Tyser AR, Beckmann J, Franklin JD, et al. Evaluation of the PROMIS physical function computer adaptive test in the upper extremity. *J Hand Surg Am*. 2014;39(10):2047–2051.e4.
- Overbeek CL, Nota SP, Jayakumar P, Hageman MG, Ring D. The PROMIS physical function correlates with the *QuickDASH* in patients with upper extremity illness. *Clin Orthop Relat Res*. 2015;473(1):311–317.
- Juniper EF, Guyatt GH, Willan A, Griffith LE. Determining a minimal important change in a disease-specific Quality of Life Questionnaire. *J Clin Epidemiol*. 1994;47(1):81–87.
- Wright A, Hannon J, Hegedus EJ, Kavchak AE. Clinimetrics corner: a closer look at the minimal clinically important difference (MCID). *J Man Manip Ther*. 2012;20(3):160–166.
- Hung M, Saltzman CL, Kendall R, et al. What are the MCIDs for PROMIS, NDI, and ODI instruments among patients with spinal conditions? *Clin Orthop Relat Res*. 2018;476(10):2027–2036.
- McGlothlin AE, Lewis RJ. Minimal clinically important difference: defining what really matters to patients. *JAMA*. 2014;312(13):1342–1343.
- Revicki D, Hays RD, Cella D, Sloan J. Recommended methods for determining responsiveness and minimally important differences for patient-reported outcomes. *J Clin Epidemiol*. 2008;61(2):102–109.
- Bernstein DN, Houck JR, Gonzalez RM, et al. Preoperative PROMIS scores predict postoperative PROMIS score improvement for patients undergoing hand surgery. *Hand (N Y)*. 2020;15(2):185–193.
- Bernstein DN, Houck JR, Mahmood B, Hammert WC. Minimal clinically important differences for PROMIS physical function, upper extremity, and pain interference in carpal tunnel release using region- and condition-specific PROM tools. *J Hand Surg Am*. 2019;44(8):635–640.
- Sandvall B, Okoroafor UC, Gerull W, Guattery J, Calfee RP. Minimal clinically important difference for PROMIS physical function in patients with distal radius fractures. *J Hand Surg Am*. 2019;44(6):454–459.e1.
- Lee DJ, Calfee RP. The minimal clinically important difference for PROMIS physical function in patients with thumb carpometacarpal arthritis. *Hand (N Y)*. 2021;16(5):638–643.
- Kazmers NH, Qiu Y, Yoo M, Stephens AR, Tyser AR, Zhang Y. The minimal clinically important difference of the PROMIS and *QuickDASH* instruments in a nonshoulder hand and upper extremity patient population. *J Hand Surg Am*. 2020;45(5):399–407.e6.
- HealthMeasures. What are the differences between versions of PROMIS physical function? Accessed January 4, 2021. <https://www.healthmeasures.net/forum-healthmeasures/promis/125-what-are-the-differences-between-versions-of-promis-physical-function>
- HealthMeasures. PROMIS scoring manuals. Accessed January 4, 2021. <https://www.healthmeasures.net/promis-scoring-manuals>
- HealthMeasures. PROMIS reference populations. Accessed May 1, 2021. https://www.healthmeasures.net/index.php?option=com_content&view=category&layout=blog&id=235&Itemid=1227
- HealthMeasures. PROMIS. Accessed January 4, 2021. <https://www.healthmeasures.net/score-and-interpret/interpret-scores/promis>
- Yost KJ, Eton DT, Garcia SF, Cella D. Minimally important differences were estimated for six Patient-Reported Outcomes Measurement Information System—cancer scales in advanced-stage cancer patients. *J Clin Epidemiol*. 2011;64(5):507–516.
- Yost KJ, Cella D, Chawla A, et al. Minimally important differences were estimated for the Functional Assessment of Cancer

- Therapy—Colorectal (FACT-C) instrument using a combination of distribution- and anchor-based approaches. *J Clin Epidemiol*. 2005;58(12):1241–1251.
37. Yost KJ, Eton DT. Combining distribution- and anchor-based approaches to determine minimally important differences: the FACIT experience. *Eval Health Prof*. 2005;28(2):172–191.
 38. Malay S, SUN Study Group, Chung KC. The minimal clinically important difference after simple decompression for ulnar neuropathy at the elbow. *J Hand Surg Am*. 2013;38(4):652–659.
 39. Jones PW, Beeh KM, Chapman KR, Decramer M, Mahler DA, Wedzicha JA. Minimal clinically important differences in pharmacological trials. *Am J Respir Crit Care Med*. 2014;189(3):250–255.
 40. Ostelo RW, Deyo RA, Stratford P, et al. Interpreting change scores for pain and functional status in low back pain: towards international consensus regarding minimal important change. *Spine (Phila Pa 1976)*. 2008;33(1):90–94.
 41. Terwee CB, Roorda LD, Dekker J, et al. Mind the MIC: large variation among populations and methods. *J Clin Epidemiol*. 2010;63(5):524–534.
 42. Ho B, Houck JR, Flemister AS, et al. Preoperative PROMIS scores predict postoperative success in foot and ankle patients. *Foot Ankle Int*. 2016;37(9):911–918.
 43. Kovacs FM, Abaira V, Royuela A, et al. Minimum detectable and minimal clinically important changes for pain in patients with nonspecific neck pain. *BMC Musculoskelet Disord*. 2008;9:43.
 44. Shauver MJ, Chung KC. The minimal clinically important difference of the Michigan hand outcomes questionnaire. *J Hand Surg Am*. 2009;34(3):509–514.
 45. Wang YC, Hart DL, Stratford PW, Mioduski JE. Baseline dependency of minimal clinically important improvement. *Phys Ther*. 2011;91(5):675–688.
 46. Coretti S, Ruggeri M, McNamee P. The minimum clinically important difference for EQ-5D index: a critical review. *Expert Rev Pharmacoecon Outcomes Res*. 2014;14(2):221–233.
 47. Sorensen AA, Howard D, Tan WH, Ketchersid J, Calfee RP. Minimal clinically important differences of 3 patient-rated outcomes instruments. *J Hand Surg Am*. 2013;38(4):641–649.